



Article

# Polygenic Score for Body Mass Index Is Associated with Weight Loss and Lipid Outcomes After Metabolic and Bariatric Surgery

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## Abstract

Metabolic and bariatric surgery (MBS) is an effective treatment for severe obesity, though individual responses vary widely, partly due to genetic predisposition. This study investigates the association of a body mass index (BMI) polygenic score (PGS) with weight loss and metabolic outcomes following surgery. A cohort of 225 patients undergoing MBS was analyzed at baseline ( $T_0$ ), six ( $T_6$ ), and twelve ( $T_{12}$ ) months, with anthropometric and biochemical parameters recorded at each time point. Total weight loss (TWL) and excess weight loss (EWL) percentages were calculated. PGS was computed using the LDpred-grid Bayesian method. The mean age was  $45.9 \pm 9.4$  years. Males had a higher baseline prevalence of type 2 diabetes (T2D) and comorbidities ( $p < 0.001$ ). Linear regression analysis confirmed an association between PGS and baseline BMI ( $p = 0.012$ ). Moreover, mediation analysis revealed that baseline BMI mediated the effect of the PGS on %TWL at  $T_{12}$ , with an indirect effect ( $p$ -value = 0.018). In contrast, high-density lipoprotein-cholesterol (HDL-C) at  $T_6$  and triglycerides (TG) at  $T_{12}$  showed direct associations with the PGS ( $p$ -value = 0.004 and  $p$ -value = 0.08, respectively), with no significant mediation by BMI. This study showed a BMI-mediated association of PGS with %TWL and a direct association with lipid changes, suggesting its potential integration into personalized obesity treatment.

**Keywords:** obesity; metabolic and bariatric surgery; polygenic score



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## 1. Introduction

Obesity is a multifactorial disease influenced by a complex interplay of behavioral, psychosocial, environmental factors and genetic predisposition [1]. According to the World Health Organization (WHO), in 2022, one in eight people worldwide were obese, and 2.5 billion adults ( $\geq 18$  years) were overweight. Moreover, by 2024, 35 million children under five years of age were classified as overweight, highlighting the early onset and widespread nature of this condition [2].

Of all available treatment options, metabolic and bariatric surgery (MBS) is currently the most effective intervention for severe obesity, leading to substantial and sustained

weight loss, as well as significant improvements in obesity-related comorbidities, such as type 2 diabetes (T2D), cardiovascular disease, obstructive sleep apnea, dyslipidemia, and hypertension [3–5].

Among the surgical techniques, sleeve gastrectomy (SG) and Roux-en-Y gastric bypass (RYGB) represent the two most performed bariatric procedures, each inducing substantial weight loss through different mechanisms. SG reduces stomach volume, alters gut hormones, such as ghrelin and peptide YY, and changes gastric motility [6]. Conversely, RYGB involves the creation of a small gastric pouch and rerouting the small intestine to this pouch, thereby bypassing much of the stomach and duodenum [7].

Although RYGB is considered the gold standard after restrictive surgery, one anastomosis gastric bypass (OAGB), also known as a mini gastric bypass, is a viable alternative due to its technical simplicity, reversibility, and shorter operating time [8]. The procedure involves suturing the stomach from its lower part to create a long gastric pouch; the rest of the stomach remains inside, but food does not enter [9]. While OAGB involves creating a single connection between the stomach and the small intestine, RYGB involves an additional connection between two sections of the small intestine [10].

Despite the overall success of these procedures, post-operative weight loss outcomes vary widely among individuals [11,12]. This inter-individual variability represents a significant clinical challenge and underscores the need to improve our knowledge of the factors that influence surgical efficacy. Understanding these determinants is essential to optimize patient selection and post-operative management.

Genetic factors have increasingly been recognized as important contributors to both obesity susceptibility and response to weight loss interventions [13]. Inherited predisposition to obesity can result from both monogenic and polygenic mechanisms. In rare cases, obesity may be the result of single-gene mutations that exert large effects on energy balance or adipose tissue regulation. However, for most individuals, no single causative mutation has been identified and genetic susceptibility is explained by the combined influence of numerous genetic variants, each exerting a modest effect on body weight regulation [14]. Indeed, in recent years, polygenic scores (PGSs), which combine the effects of numerous genetic variants carried by an individual associated with a given trait or disease, have emerged as promising tools for quantifying inherited predisposition to obesity [15,16].

Several studies have demonstrated that individuals with a higher PGS for body mass index (BMI) or obesity are more likely to gain weight over time and have a greater risk of developing obesity-related diseases [17,18]. However, the role of polygenic risk in predicting weight loss following MBS remains an area of active investigation. While some evidence suggests that a greater genetic predisposition to obesity may be associated with reduced post-surgical weight loss or greater weight regain over time after surgery [19–21], other studies report no significant association [22]. These inconsistencies may arise from differences in PGS construction, sample sizes, ethnic backgrounds of study populations, or types of surgical procedure analyzed.

Given the growing interest in precision medicine approaches to metabolic surgery, clarifying the relationship between obesity-related polygenic risk and weight loss trajectories after MBS is of both clinical and scientific importance. Therefore, this study aims to assess whether a BMI-based PGS is associated with weight loss outcomes following bariatric surgery, measured by total weight loss (TWL). Moreover, the association of PGS with other metabolic outcomes, such as lipid profile and hepatic enzyme levels, was evaluated to better understand the broader physiological impact of genetic risk in the context of metabolic surgery.

## 2. Results

### 2.1. Sample Characteristics

The study cohort, composed mainly of females (73.8%), had a mean age of  $45.9 \pm 9.4$ , ranging from 20 to 63. Males showed a higher percentage of T2D at T<sub>0</sub> and other comorbidities compared to females ( $p$ -value < 0.001) (Table 1).

**Table 1.** Sample characteristics at enrollment by gender.  $p$ -value from Wilcoxon test and Chi-squared test for continuous and categorical variables, respectively. RYGB, Roux-en-Y gastric bypass; OAGB, one anastomosis gastric bypass; SG, sleeve gastrectomy.

Characteristics	All ( $n = 225$ )	Males ( $n = 59$ )	Females ( $n = 166$ )	$p$ -Value
Age (mean $\pm$ SD)	$45.9 \pm 9.4$	$47.0 \pm 9.1$	$45.5 \pm 9.5$	0.24
Surgery (%)				
RYGB	39.1	36.0	40.4	0.61
OAGB	27.1	25.0	27.7	
SG	33.8	39.0	31.9	
Diabetes (% Yes)	27.6	52.5	18.7	<0.001
Comorbidities (% Yes)	66.2	84.7	59.6	<0.001

High-density lipoprotein-cholesterol (HDL-C), low-density lipoprotein-cholesterol (LDL-C), and total cholesterol (TC) are significantly higher in females at all time points, while males had higher levels of triglycerides (TG) at T<sub>0</sub> ( $p$ -value = 0.013) (Table 2).

Regarding hepatic profile, alanine aminotransferase (ALT), aspartate aminotransferase (AST) and gamma-glutamyl transferase (GGT) were significantly different between males and females at T<sub>0</sub> ( $p$ -value < 0.001). Differences in hepatic enzymes were also found for AST at T<sub>6</sub> ( $p$ -value = 0.018) and T<sub>12</sub> ( $p$ -value = 0.007) and for GGT at T<sub>6</sub> ( $p$ -value < 0.001) and T<sub>12</sub> ( $p$ -value < 0.001) (Table 2).

Compared to RYGB and OAGB patients, those who underwent SG had a significantly higher BMI at T<sub>0</sub> ( $p$ -value<sub>RYGB</sub> = 0.006,  $p$ -value<sub>OAGB</sub> < 0.001), at T<sub>6</sub> ( $p$ -value<sub>RYGB</sub> = 0.006,  $p$ -value<sub>OAGB</sub> < 0.001), and at T<sub>12</sub> ( $p$ -value<sub>RYGB</sub> < 0.001,  $p$ -value<sub>OAGB</sub> < 0.001), while no differences were found when comparing RYGB and OAGB patients with each other (Figure S1). Obese subjects who underwent SG had a significant lower %TWL compared to those who underwent RYGB at T<sub>6</sub> ( $p$ -value = 0.012), but not at T<sub>12</sub> (Figure S2).

**Table 2.** Anthropometric, lipids, and hepatic enzymes at T<sub>0</sub>, T<sub>6</sub>, and T<sub>12</sub>.  $p$ -value from  $t$ -test or Wilcoxon test for normally and not normally distributed data, respectively. BMI, body mass index; %TWL, percentage of total weight loss; %EWL, percentage of excess weight loss; IQR, interquartile range; HDL-C, high-density lipoprotein-cholesterol; LDL-C, low-density lipoprotein-cholesterol; TC, total cholesterol; TG, triglycerides; ALT, alanine aminotransferase; AST, aspartate aminotransferase; GGT, gamma-glutamyl transferase; HbA1c, glycated hemoglobin.

Characteristics	All	Males	Females	$p$ -Value
BMI, kg/m <sup>2</sup> , mean $\pm$ SD				
T <sub>0</sub>	$42.8 \pm 5.5$	$43.6 \pm 5.5$	$42.5 \pm 5.5$	0.19
T <sub>6</sub>	$31.5 \pm 4.3$	$32 \pm 4.3$	$31.3 \pm 4.3$	0.33
T <sub>12</sub>	$28.8 \pm 4.4$	$29.4 \pm 3.6$	$28.5 \pm 4.6$	0.13
TWL, %, mean $\pm$ SD				
T <sub>6</sub>	$26.2 \pm 6.2$	$26.3 \pm 5.6$	$26.1 \pm 6.3$	0.45
T <sub>12</sub>	$32.4 \pm 8.5$	$31.7 \pm 8.1$	$32.6 \pm 8.7$	0.49
EWL, %, mean $\pm$ SD				
T <sub>6</sub>	$66.2 \pm 18.5$	$65.1 \pm 19.0$	$66.6 \pm 18.3$	0.47
T <sub>12</sub>	$81.7 \pm 23.1$	$77.7 \pm 20.0$	$83.0 \pm 24.1$	0.46

Table 2. Cont.

Characteristics	All	Males	Females	p-Value
HDL-C, mg/dL, median (IQR)				
T <sub>0</sub>	48 (15)	40 (11.5)	50 (13)	<0.001
T <sub>6</sub>	50 (17)	44 (14)	53 (16)	0.002
T <sub>12</sub>	57 (17)	50.5 (18.5)	60 (16.2)	<0.001
LDL-C, mg/dL, median (IQR)				
T <sub>0</sub>	133 (43.6)	114.4 (53)	140.2 (41.3)	<0.001
T <sub>6</sub>	102.2 (42)	92.6 (45.9)	104 (38.3)	0.03
T <sub>12</sub>	98.6 (47.9)	84.4 (47)	101.2 (45)	0.002
TC, mg/dL, median (IQR)				
T <sub>0</sub>	209 (60)	180 (60)	215 (51)	<0.001
T <sub>6</sub>	174 (49.5)	147 (56)	178.5 (43)	0.003
T <sub>12</sub>	175 (52.5)	151 (50)	181 (48)	<0.001
TG, mg/dL, median (IQR)				
T <sub>0</sub>	117 (66)	151 (104.5)	113 (55.5)	0.013
T <sub>6</sub>	91 (39)	87 (50.5)	93 (35.8)	0.48
T <sub>12</sub>	83 (39.5)	79 (34)	84 (41)	0.83
ALT, U/L, median (IQR)				
T <sub>0</sub>	24 (17)	32 (23)	21 (13.8)	<0.001
T <sub>6</sub>	18 (10)	21 (16)	17 (8)	0.08
T <sub>12</sub>	17 (13.3)	21 (17.5)	16 (12)	0.07
AST, U/L, median (IQR)				
T <sub>0</sub>	21 (10)	26 (14.5)	20 (8)	<0.001
T <sub>6</sub>	19 (8)	21 (7.8)	18 (7)	0.018
T <sub>12</sub>	19 (9)	22 (10)	18 (7)	0.007
GGT, U/L, median (IQR)				
T <sub>0</sub>	24 (25)	39 (33)	21.5 (17.8)	<0.001
T <sub>6</sub>	15 (12.8)	21.5 (18.8)	12 (11)	<0.001
T <sub>12</sub>	15 (11.5)	21 (20)	13 (8.8)	<0.001
HbA1c, %, median (IQR)				
T <sub>0</sub>	5.8 (0.7)	6.2 (1.2)	5.8 (0.5)	0.014
T <sub>6</sub>	5.4 (0.6)	5.4 (0.6)	5.4 (0.6)	0.90
T <sub>12</sub>	5.3 (0.6)	5.2 (0.7)	5.4 (0.5)	0.24

## 2.2. Association of PGS with Weight Loss and Metabolic Outcomes

Linear regression analysis with age and sex as covariates confirmed in our cohort that PGS is associated with BMI at T<sub>0</sub> ( $p$ -value = 0.012,  $\beta$  =  $-4.7$ ,  $R^2$  = 5%).

Therefore, to study the effect of PGS on weight loss and metabolic outcomes, mediation analysis where BMI at T<sub>0</sub> was considered as the mediator was conducted. This analysis showed that baseline BMI significantly mediated the association between the PGS and post-operative outcomes (Table 3). Specifically, for %TWL at T<sub>12</sub>, the indirect effect was significant ( $p$ -value = 0.018), while the direct effect did not reach statistical significance ( $p$ -value = 0.144). The total effect was statistically significant ( $p$ -value = 0.028), and approximately 39% of the total effect was mediated by baseline BMI ( $p$ -value = 0.046). Similarly, for %TWL at T<sub>6</sub>, the indirect effect was also significant ( $p$ -value = 0.014), but neither the direct nor the total effect reached statistical significance.

Regarding %EWL at T<sub>6</sub> and T<sub>12</sub>, the average casual mediation effect (ACME) was significant ( $p$ -value = 0.024 and  $p$ -value = 0.012, respectively), whereas the direct and total effects were not statistically significant.

Mediation analysis for lipids indicated a significant total effect of the PGS on variation of HDL-C at T<sub>6</sub> ( $p$ -value < 0.001) and TG at T<sub>12</sub> ( $p$ -value = 0.026) post-surgery. The direct effect was significant for both HDL-C at T<sub>6</sub> ( $p$ -value = 0.004) and TG at T<sub>12</sub> ( $p$ -value = 0.008), while the indirect effect mediated by baseline BMI did not reach statistical significance.

**Table 3.** Results of mediation analysis for weigh loss outcomes. Sex, age, and type of surgery have been used as covariates. %TWL, percentage of total weight loss; % EWL, percentage of excess weight loss. ACME, Average Causal Mediation Effect; ADE, Average Direct Effect; CI, confidence interval.

Variable	Estimate	p-Value	95% CI Lower	95% CI Upper
<b>%TWL T<sub>6</sub></b>				
ACME	−1.22	0.014	−2.56	−0.19
ADE	−1.07	0.542	−4.97	3.03
Total effect	−2.29	0.278	−6.23	1.76
Proportion mediated	0.53	0.292	−3.61	3.50
<b>%TWL T<sub>12</sub></b>				
ACME	−2.62	0.018	−5.08	−0.53
ADE	−4.06	0.144	−9.70	1.43
Total effect	−6.68	0.028	−12.47	−1.09
Proportion mediated	0.39	0.046	0.030	1.48
<b>%EWL T<sub>6</sub></b>				
ACME	7.95	0.024	1.27	14.58
ADE	−3.73	0.482	−13.62	6.42
Total effect	4.22	0.524	−7.65	16.31
Proportion mediated	1.88	0.512	−18.15	11.77
<b>%EWL T<sub>12</sub></b>				
ACME	7.31	0.012	1.55	14.47
ADE	−10.94	0.170	−24.26	4.75
Total effect	−3.63	0.674	−18.12	12.67
Proportion mediated	−2.01	0.686	−9.95	10.85

The proportion mediated was not significant, suggesting that the effect of the PGS on triglyceride changes is primarily direct, rather than mediated by baseline BMI (Table 4).

**Table 4.** Results of mediation analysis for lipids Δ values. Sex, age, and type of surgery have been used as covariates. Δ values for HDL-C, LDL-C, TG and TC were obtained as the difference between the values at T<sub>6</sub> and T<sub>0</sub> for ΔT<sub>6</sub> and the difference between the values at T<sub>12</sub> and T<sub>0</sub> for ΔT<sub>12</sub>. HDL-C, high-density lipoprotein-cholesterol; LDL-C, low-density lipoprotein-cholesterol; TG, triglycerides; TC, total cholesterol; ACME, Average Causal Mediation Effect; ADE, Average Direct Effect, CI, confidence interval.

Variable	Estimate	p-Value	95% CI Lower	95% CI Upper
<b>ΔHDL-C T<sub>6</sub></b>				
ACME	0.00007	0.95	−0.046	0.04
ADE	0.3	0.004	0.062	0.6
Total effect	0.3	<0.001	0.084	0.6
Proportion mediated	0.0002	0.95	−0.1	0.3
<b>ΔHDL-C T<sub>12</sub></b>				
ACME	−0.02	0.34	−0.08	0.015
ADE	0.17	0.12	−0.03	0.46
Total effect	0.15	0.14	−0.04	0.41
Proportion mediated	−0.13	0.4	−0.8	0.78
<b>ΔLDL-C T<sub>6</sub></b>				
ACME	−0.01	0.78	−0.096	0.037
ADE	0.4	0.076	−0.029	0.998
Total effect	0.4	0.058	−0.007	0.95
Proportion mediated	−0.028	0.73	−0.28	0.34

**Table 4.** *Cont.*

Variable	Estimate	<i>p</i> -Value	95% CI Lower	95% CI Upper
$\Delta$ LDL-C T <sub>12</sub>				
ACME	0.024	0.28	−0.016	0.08
ADE	−0.073	0.59	−0.32	0.18
Total effect	−0.048	0.73	−0.29	0.2
Proportion mediated	−0.5	0.82	−3.54	2.5
$\Delta$ TG T <sub>6</sub>				
ACME	−0.003	0.89	−0.057	0.048
ADE	−0.19	0.21	−0.47	0.11
Total effect	−0.2	0.18	−0.47	0.09
Proportion mediated	0.016	0.98	−1.1	0.9
$\Delta$ TG T <sub>12</sub>				
ACME	0.056	0.082	−0.0056	0.15
ADE	−0.34	0.008	−0.6	−0.096
Total effect	−0.3	0.026	−0.55	−0.04
Proportion mediated	−0.18	0.108	−1.15	0.034
$\Delta$ TC T <sub>6</sub>				
ACME	0.008	0.54	−0.02	0.04
ADE	0.068	0.48	−0.09	0.25
Total effect	0.075	0.37	−0.08	0.25
Proportion mediated	0.1	0.8	−1.7	1.88
$\Delta$ TC T <sub>12</sub>				
ACME	−0.004	0.90	−0.063	0.038
ADE	0.025	0.78	−0.16	0.25
Total effect	0.02	0.79	−0.16	0.21
Proportion mediated	−0.2	0.81	−3.3	2.75

No significant associations with hepatic enzymes (ALT, AST, and GGT) and HbA1c (glycated hemoglobin) have emerged (Table S1).

### 3. Discussion

This study investigated the role of a PGS for BMI in influencing weight loss and metabolic outcomes following MBS.

Our findings show that genetic predisposition to higher BMI, as captured by the PGS, has a significant indirect influence on weight loss (measured as %TWL) at twelve months after surgery, with baseline BMI acting as a key mediator. Specifically, nearly 40% of the effect of the PGS on %TWL was mediated by baseline BMI, highlighting the role of genetic predisposition in shaping pre-operative body composition, which, in turn, influences the magnitude and variability of weight loss in response to MBS.

In this study, we also identified a direct association of the PGS with changes in lipid profile after bariatric surgery, specifically changes in HDL-C at six months and TG levels at twelve months postoperatively. Unlike %TWL, these associations with lipids were not mediated by baseline BMI, suggesting that genetic predisposition to obesity may influence the metabolic response to surgery, particularly with respect to lipid metabolism. This aligns with the concept of “metabolic responders” versus “non-responders” to MBS, suggesting that genomic factors may contribute to such variability [23].

In fact, MBS is increasingly recognized not only as a weight reduction strategy, but as a metabolic intervention with far-reaching effects on cardiovascular and endocrine health. Accordingly, current definitions of surgical success now emphasize improvements in metabolic outcomes, such as glycemic control, lipid levels, and cardiometabolic risk

reduction, rather than weight loss alone [24–26]. Lipid profile changes, including increases in HDL-C and reductions in TG, are among the most clinically meaningful metabolic shifts observed after surgery and are strong predictors of cardiovascular risk reduction [27,28].

Moreover, our findings agree with prior research indicating shared genetic determinants between obesity and lipid traits [29,30].

The results of this study are also consistent with previous literature highlighting sex-specific metabolic responses and support the need for gender-tailored clinical management in patients with obesity [31,32]. In our cohort, sex differences emerged in several clinical variables: males exhibited both higher TG levels and prevalence of T2D and other comorbidities at baseline, as well as elevated hepatic enzyme levels throughout follow-up, whereas females consistently presented higher levels of HDL-C, LDL-C, and TC. Moreover, regarding the type of surgery, our study showed that patients who underwent SG had a significantly higher BMI at all time points compared to those who underwent RYGB or OAGB and showed a lower %TWL at six months after surgery compared to RYGB patients, confirming previous studies that have already reported that the type of bariatric surgery influences the weight loss trajectory [33,34].

The use of mediation analysis to investigate the complex interplay between genetic predisposition, baseline BMI, and both anthropometric and metabolic outcomes following bariatric surgery is a strength of this study. This method allowed us to distinguish direct from indirect genetic effects and to uncover the mediating role of pre-operative BMI in influencing weight loss outcomes.

This study has limitations, however. The sample size, while adequate for initial exploration, may not capture the full spectrum of genetic variability, and replication in larger, independent cohorts is necessary. Furthermore, the observed effect of PGS on weight loss outcomes was modest in magnitude and largely mediated by baseline BMI. However, given the polygenic architecture of obesity and the multifactorial nature of post-surgical weight loss, this result is expected, suggesting that environmental and behavioral factors likely play a substantial role and should be integrated into future models. Importantly, our results highlight that BMI-PGS should not be considered a replacement for established clinical predictors, such as baseline BMI, but rather a complementary source of information that may improve risk stratification or identify individuals who may benefit from more intensive follow-up.

Therefore, as genotyping becomes more accessible in clinical settings, although PGS cannot yet inform individual clinical decisions, integrating genetic data into clinical decision-making, alongside demographic, clinical, and behavioral factors, could enhance personalized approaches to obesity treatment, ultimately improving patient outcomes after bariatric surgery.

## 4. Materials and Methods

### 4.1. Participants

A total of 225 severely obese subjects undergoing MBS were enrolled and followed up for one year. The inclusion criteria were: well-informed and motivated patients with acceptable surgical risk; failure of previous non-surgical weight loss treatments; aged between 18 and 65 years old; and a BMI (body mass index,  $\text{kg}/\text{m}^2$ )  $> 40 \text{ kg}/\text{m}^2$  or  $>35 \text{ kg}/\text{m}^2$ , with comorbidities related to obesity. Surgical interventions consisted of RYGB (39.1%), OAGB (27.1%), and SG (33.8%). Patients were evaluated at three time points: before surgery ( $T_0$ ) and at six ( $T_6$ ) and twelve ( $T_{12}$ ) months after surgery. All subjects gave their written informed consent before participating in this study, approved by protocol N. 16,637 Local Ethical Committee (Comitato Etico Unico Regionale FVG, Udine, Italy).

#### 4.2. Data

Demographic and clinical data were collected for each participant at all three time points. Weight (W) and height were taken on a digital weighing scale, and BMI was calculated ( $\text{kg}/\text{m}^2$ ).

The total weight loss (%TWL) at  $T_6$  and  $T_{12}$  was calculated as  $(\text{Weight loss}/WT_0) \times 100$ .

The excess weight loss (%EWL) at  $T_6$  and  $T_{12}$  was calculated as  $(\text{Weight loss}/\text{Excess Weight}) \times 100$  (where excess weight is  $T_0$  weight – ideal weight).

Data on diabetes diagnosis, ongoing diabetic therapy, and the presence of other comorbidities were also recorded. Moreover, lipid profile (HDL-C, LDL-C, TC and TG), hepatic enzymes (ALT, AST and GGT), and HbA1c were measured.

Delta values for HDL-C, LDL-C, TC, TG, ALT, AST, GGT, and HbA1c were calculated as the difference between the variable at  $T_6$  and  $T_0$  for  $\Delta T_6$  and between the variable at  $T_{12}$  and  $T_0$  for  $\Delta T_{12}$ .

#### 4.3. DNA Extraction and Genotyping

For each participant,  $23 \pm 9$  mg of liver tissue was collected and stored at  $-80^\circ\text{C}$  until DNA extraction, which was performed with the Quick-DNA/RNA™ Miniprep Plus Kit (Zymo Research, Irvine, CA, USA), following the manufacturer's protocol. Frozen samples were homogenized in 300  $\mu\text{L}$  of DNA/RNA Shield™ (1X) using a bead homogenizer (Bead Ruptor, Omni Inc, Kennesaw, GA, USA), then incubated with 15  $\mu\text{L}$  of Proteinase K and 30  $\mu\text{L}$  of PK Digestion Buffer at room temperature (RT) for 30 min. DNA/RNA Lysis Buffer was added in a 1:1 ratio, and the lysates were processed through Spin-Away™ Filters and purification columns, following the protocol's instruction (Zymo Research, Irvine, CA, USA). DNA was then eluted in 50  $\mu\text{L}$  of DNase/RNase-free water. Subsequently, the concentration ( $\text{ng}/\mu\text{L}$ ) of the nucleic acid was determined using the NanoDrop® 1000 spectrophotometer (Thermo Fisher Scientific, Waltham, MA, USA). The absorbance ratios at 260/280 and 260/230 were evaluated as indicators of the DNA purity. Then, 96-well plates were prepared with DNA samples adjusted to a final concentration of 65  $\text{ng}/\mu\text{L}$  in a total volume of 10  $\mu\text{L}$  per well. Genotyping was conducted by Illumina Infinium Global Screening Array (GSA v3.0).

#### 4.4. Quality Control, Imputation, and PGS Calculation

Quality control (QC) was performed on genotyped data to remove samples with low genotype calls (<95%), heterozygosity with six standard deviations (SDs) away from the mean, identity by descent (IBD) proportion > 0.4, and to remove duplicate single nucleotide polymorphisms (SNPs) and those with missing call rate > 1% or with a Hardy–Weinberg equilibrium (HWE)  $p$ -value <  $1 \times 10^{-6}$ . Moreover, the PLINKv1.9—check-sex function made it possible to assess the potential sex discrepancies between the predicted and reported sex of individuals [35,36]. Samples with a non-European ancestry detected by principal component analysis were excluded. After the QC procedure, the genotyped data were subjected to a genotyping imputation using standard procedures. Briefly, genotype phasing was performed with the software Eagle2 V2.4.1 [37], while impute5 V1.1.5 software [38] and IGRP panel [39] were used for imputation. Variants with imputation scores < 0.8 were removed.

The dedicated LDpred2 program using a Bayesian approach was used to calculate PGS [40]. Imputed data as target data and summary statistics from a published BMI Genome-Wide Association Study (GWAS) [41] as base data were used. SNPs from the BMI GWAS were matched with SNPs from our imputed data and only SNPs that were present in both were considered ( $n = 871,416$ ). For each SNP, the number of the alleles (0, 1, or 2)

associated with BMI was multiplied by the effect estimated in the GWAS. The PGS for each individual was an average of the weighted BMI-associated alleles.

#### 4.5. Statistical Analyses

Continuous variables are described as mean  $\pm$  SD when normally distributed, and as median (IQR) when not normally distributed. Categorical variables are reported as percentages. Differences between male and female were assessed using the *t*-test or Wilcoxon rank sum test for continuous variables that were normally or not normally distributed, respectively; while the Chi-squared test was used for categorical variables. Due to their not normal distributions, lipids and hepatic enzymes data were log-transformed prior to calculating the  $\Delta$  values at T<sub>6</sub> and T<sub>12</sub>. To study the effect of PGS on post-surgical weight loss and metabolic outcomes, we conducted a causal mediation analysis using the `mediate()` function from the R mediation package (V4.4.2). The average causal mediation effect (ACME), average direct effect (ADE), total effect, and proportion mediated were estimated using nonparametric bootstrapping with 1000 simulations. The PGS was treated as the exposure, baseline BMI as the mediator, and post-operative weight loss and metabolic outcomes as the dependent variables. The analyses were adjusted for age, sex, and type of surgery.

Linear regression was conducted to test the association between PGS and baseline BMI, with correction for age and sex. ANOVA and the post-hoc Tukey's Honest Significant Difference (HSD) test were used to study the relationship between %TWL and type of surgery and between BMI and type of surgery. All statistical analyses were conducted in RStudio (V4.4.2) [42].

**Supplementary Materials:** The following supporting information can be downloaded at: <https://www.mdpi.com/article/10.3390/ijms26157337/s1>.

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## References

1. Maxim, M.; Soroceanu, R.P.; Vlăsceanu, V.I.; Platon, R.L.; Toader, M.; Miler, A.A.; Onofriescu, A.; Abdulan, I.M.; Ciuntu, B.-M.; Balan, G.; et al. Dietary Habits, Obesity, and Bariatric Surgery: A Review of Impact and Interventions. *Nutrients* **2025**, *17*, 474. [CrossRef] [PubMed]
2. Obesity and Overweight. Available online: <https://www.who.int/news-room/fact-sheets/detail/obesity-and-overweight> (accessed on 28 May 2025).
3. Busebee, B.; Ghusn, W.; Cifuentes, L.; Acosta, A. Obesity: A Review of Pathophysiology and Classification. *Mayo Clin. Proc.* **2023**, *98*, 1842–1857. [CrossRef] [PubMed]
4. Sjöström, L.; Narbro, K.; Sjöström, C.D.; Karason, K.; Larsson, B.; Wedel, H.; Lystig, T.; Sullivan, M.; Bouchard, C.; Carlsson, B.; et al. Effects of Bariatric Surgery on Mortality in Swedish Obese Subjects. *N. Engl. J. Med.* **2007**, *357*, 741–752. [CrossRef] [PubMed]
5. Adams, T.D.; Davidson, L.E.; Litwin, S.E.; Kim, J.; Kolotkin, R.L.; Nanjee, M.N.; Gutierrez, J.M.; Frogley, S.J.; Ibele, A.R.; Brinton, E.A.; et al. Weight and Metabolic Outcomes 12 Years after Gastric Bypass. *N. Engl. J. Med.* **2017**, *377*, 1143–1155. [CrossRef]
6. Aderinto, N.; Olatunji, G.; Kokori, E.; Olaniyi, P.; Isarinade, T.; Yusuf, I.A. Recent Advances in Bariatric Surgery: A Narrative Review of Weight Loss Procedures. *Ann. Med. Surg.* **2023**, *85*, 6091–6104. [CrossRef]
7. Mitchell, B.G.; Collier, S.A.; Gupta, N. Roux-En-Y Gastric Bypass. In *StatPearls*; StatPearls Publishing: Treasure Island, FL, USA, 2025.
8. Poublon, N.; Chidi, I.; Bethlehem, M.; Kuipers, E.; Gadiot, R.; Emous, M.; van Det, M.; Dunkelgrun, M.; Biter, U.; Apers, J. One Anastomosis Gastric Bypass vs. Roux-En-Y Gastric Bypass, Remedy for Insufficient Weight Loss and Weight Regain after Failed Restrictive Bariatric Surgery. *Obes. Surg.* **2020**, *30*, 3287–3294. [CrossRef]
9. Arakkakunnel, J.; Grover, K. One Anastomosis Gastric Bypass and Mini Gastric Bypass. In *StatPearls*; StatPearls Publishing: Treasure Island, FL, USA, 2025.
10. Robert, M.; Poghosyan, T.; Maucourt-Boulch, D.; Filippello, A.; Caiazzo, R.; Sterkers, A.; Khamphommala, L.; Reche, F.; Malherbe, V.; Torcivia, A.; et al. Efficacy and Safety of One Anastomosis Gastric Bypass versus Roux-En-Y Gastric Bypass at 5 Years (YOMEGA): A Prospective, Open-Label, Non-Inferiority, Randomised Extension Study. *Lancet Diabetes Endocrinol.* **2024**, *12*, 267–276. [CrossRef]
11. Courcoulas, A.P.; Goodpaster, B.H.; Eagleton, J.K.; Belle, S.H.; Kalarchian, M.A.; Lang, W.; Toledo, F.G.S.; Jakicic, J.M. Surgical vs Medical Treatments for Type 2 Diabetes Mellitus: A Randomized Clinical Trial. *JAMA Surg.* **2014**, *149*, 707–715. [CrossRef]
12. Arterburn, D.E.; Telem, D.A.; Kushner, R.F.; Courcoulas, A.P. Benefits and Risks of Bariatric Surgery in Adults: A Review. *JAMA* **2020**, *324*, 879–887. [CrossRef]
13. Loos, R.J.F.; Yeo, G.S.H. The Genetics of Obesity: From Discovery to Biology. *Nat. Rev. Genet.* **2022**, *23*, 120–133. [CrossRef]
14. Khera, A.V.; Chaffin, M.; Wade, K.H.; Zahid, S.; Brancale, J.; Xia, R.; Distefano, M.; Senol-Cosar, O.; Haas, M.E.; Bick, A.; et al. Polygenic Prediction of Weight and Obesity Trajectories from Birth to Adulthood. *Cell* **2019**, *177*, 587–596.e9. [CrossRef] [PubMed]
15. Khera, A.V.; Chaffin, M.; Aragam, K.G.; Haas, M.E.; Roselli, C.; Choi, S.H.; Natarajan, P.; Lander, E.S.; Lubitz, S.A.; Ellinor, P.T.; et al. Genome-Wide Polygenic Scores for Common Diseases Identify Individuals with Risk Equivalent to Monogenic Mutations. *Nat. Genet.* **2018**, *50*, 1219–1224. [CrossRef] [PubMed]
16. Locke, A.E.; Kahali, B.; Berndt, S.I.; Justice, A.E.; Pers, T.H.; Day, F.R.; Powell, C.; Vedantam, S.; Buchkovich, M.L.; Yang, J.; et al. Genetic Studies of Body Mass Index Yield New Insights for Obesity Biology. *Nature* **2015**, *518*, 197–206. [CrossRef] [PubMed]
17. Akiyama, M.; Okada, Y.; Kanai, M.; Takahashi, A.; Momozawa, Y.; Ikeda, M.; Iwata, N.; Ikegawa, S.; Hirata, M.; Matsuda, K.; et al. Genome-Wide Association Study Identifies 112 New Loci for Body Mass Index in the Japanese Population. *Nat. Genet.* **2017**, *49*, 1458–1467. [CrossRef]
18. Yengo, L.; Sidorenko, J.; Kemper, K.E.; Zheng, Z.; Wood, A.R.; Weedon, M.N.; Frayling, T.M.; Hirschhorn, J.; Yang, J.; Visscher, P.M.; et al. Meta-Analysis of Genome-Wide Association Studies for Height and Body Mass Index in ~700000 Individuals of European Ancestry. *Hum. Mol. Genet.* **2018**, *27*, 3641–3649. [CrossRef]
19. Aasbrenn, M.; Schnurr, T.M.; Have, C.T.; Svendstrup, M.; Hansen, D.L.; Worm, D.; Balslev-Harder, M.; Hollensted, M.; Grarup, N.; Burgdorf, K.S.; et al. Genetic Determinants of Weight Loss After Bariatric Surgery. *Obes. Surg.* **2019**, *29*, 2554–2561. [CrossRef]
20. Peña, E.; Mas-Bermejo, P.; Lecube, A.; Ciudin, A.; Arenas, C.; Simó, R.; Rigla, M.; Caixàs, A.; Rosa, A. Use of Polygenic Risk Scores to Assess Weight Loss after Bariatric Surgery: A 5-Year Follow-up Study. *J. Gastrointest. Surg.* **2024**, *28*, 1400–1405. [CrossRef]
21. German, J.; Cordioli, M.; Tozzo, V.; Urbut, S.; Arumãe, K.; Smit, R.A.J.; Lee, J.; Li, J.H.; Janucik, A.; Ding, Y.; et al. Association between Plausible Genetic Factors and Weight Loss from GLP1-RA and Bariatric Surgery. *Nat. Med.* **2025**, *31*, 2269–2276. [CrossRef]
22. Sarzynski, M.A.; Jacobson, P.; Rankinen, T.; Carlsson, B.; Sjöström, L.; Bouchard, C.; Carlsson, L.M.S. Associations of Markers in 11 Obesity Candidate Genes with Maximal Weight Loss and Weight Regain in the SOS Bariatric Surgery Cases. *Int. J. Obes.* **2011**, *35*, 676–683. [CrossRef]

23. Still, C.D.; Wood, G.C.; Chu, X.; Manney, C.; Strodel, W.; Petrick, A.; Gabrielsen, J.; Mirshahi, T.; Argyropoulos, G.; Seiler, J.; et al. Clinical Factors Associated with Weight Loss Outcomes after Roux-En-Y Gastric Bypass Surgery. *Obesity* **2014**, *22*, 888–894. [[CrossRef](#)]
24. Schauer, P.R.; Bhatt, D.L.; Kirwan, J.P.; Wolski, K.; Aminian, A.; Brethauer, S.A.; Navaneethan, S.D.; Singh, R.P.; Pothier, C.E.; Nissen, S.E.; et al. Bariatric Surgery versus Intensive Medical Therapy for Diabetes—5-Year Outcomes. *N. Engl. J. Med.* **2017**, *376*, 641–651. [[CrossRef](#)] [[PubMed](#)]
25. Rubino, F.; Nathan, D.M.; Eckel, R.H.; Schauer, P.R.; Alberti, K.G.M.M.; Zimmet, P.Z.; Del Prato, S.; Ji, L.; Sadikot, S.M.; Herman, W.H.; et al. Metabolic Surgery in the Treatment Algorithm for Type 2 Diabetes: A Joint Statement by International Diabetes Organizations. *Diabetes Care* **2016**, *39*, 861–877. [[CrossRef](#)] [[PubMed](#)]
26. Brethauer, S.A.; Kim, J.; el Chaar, M.; Papanavass, P.; Eisenberg, D.; Rogers, A.; Ballem, N.; Kligman, M.; Kothari, S. ASMBS Clinical Issues Committee Standardized Outcomes Reporting in Metabolic and Bariatric Surgery. *Surg. Obes. Relat. Dis.* **2015**, *11*, 489–506. [[CrossRef](#)] [[PubMed](#)]
27. Magro, D.O.; Geloneze, B.; Delfini, R.; Pareja, B.C.; Callejas, F.; Pareja, J.C. Long-Term Weight Regain after Gastric Bypass: A 5-Year Prospective Study. *Obes. Surg.* **2008**, *18*, 648–651. [[CrossRef](#)]
28. Rader, D.J.; Hovingh, G.K. HDL and Cardiovascular Disease. *Lancet* **2014**, *384*, 618–625. [[CrossRef](#)]
29. Cadby, G.; Melton, P.E.; McCarthy, N.S.; Almeida, M.; Williams-Blangero, S.; Curran, J.E.; VandeBerg, J.L.; Hui, J.; Beilby, J.; Musk, A.W.; et al. Pleiotropy of Cardiometabolic Syndrome with Obesity-Related Anthropometric Traits Determined Using Empirically Derived Kinships from the Busselton Health Study. *Hum. Genet.* **2018**, *137*, 45–53. [[CrossRef](#)]
30. Ke, J.; Gao, W.; Wang, B.; Cao, W.; Lv, J.; Yu, C.; Huang, T.; Sun, D.; Liao, C.; Pang, Y.; et al. Exploring the Genetic Association between Obesity and Serum Lipid Levels Using Bivariate Methods. *Twin Res. Hum. Genet.* **2022**, *25*, 234–244. [[CrossRef](#)]
31. Risi, R.; Rossini, G.; Tozzi, R.; Pieralice, S.; Monte, L.; Masi, D.; Castagneto-Gissey, L.; Gallo, I.F.; Strigari, L.; Casella, G.; et al. Sex Difference in the Safety and Efficacy of Bariatric Procedures: A Systematic Review and Meta-Analysis. *Surg. Obes. Relat. Dis.* **2022**, *18*, 983–996. [[CrossRef](#)]
32. Kantowski, T.; Schulze zur Wiesch, C.; Aberle, J.; Lautenbach, A. Obesity Management: Sex-Specific Considerations. *Arch. Gynecol. Obstet.* **2024**, *309*, 1745–1752. [[CrossRef](#)]
33. Lin, S.; Guan, W.; Yang, N.; Zang, Y.; Liu, R.; Liang, H. Short-Term Outcomes of Sleeve Gastrectomy plus Jejunojejunal Bypass: A Retrospective Comparative Study with Sleeve Gastrectomy and Roux-En-Y Gastric Bypass in Chinese Patients with BMI  $\geq 35$  kg/m<sup>2</sup>. *Obes. Surg.* **2019**, *29*, 1352–1359. [[CrossRef](#)]
34. Courcoulas, A.P.; King, W.C.; Belle, S.H.; Berk, P.; Flum, D.R.; Garcia, L.; Gourash, W.; Horlick, M.; Mitchell, J.E.; Pomp, A.; et al. Seven-Year Weight Trajectories and Health Outcomes in the Longitudinal Assessment of Bariatric Surgery (LABS) Study. *JAMA Surg.* **2018**, *153*, 427–434. [[CrossRef](#)]
35. Chang, C.C.; Chow, C.C.; Tellier, L.C.; Vattikuti, S.; Purcell, S.M.; Lee, J.J. Second-Generation PLINK: Rising to the Challenge of Larger and Richer Datasets. *GigaScience* **2015**, *4*, s13742-015-0047-0048. [[CrossRef](#)]
36. Purcell, S.; Neale, B.; Todd-Brown, K.; Thomas, L.; Ferreira, M.A.; Bender, D.; Maller, J.; Sklar, P.; de Bakker, P.I.; Daly, M.J.; et al. PLINK: A tool set for whole-genome association and population-based linkage analyses. *Am. J. Hum. Genet.* **2007**, *81*, 559–575. [[CrossRef](#)]
37. Loh, P.-R.; Danecek, P.; Palamara, P.F.; Fuchsberger, C.; Reshef, Y.A.; Finucane, H.K.; Schoenherr, S.; Forer, L.; McCarthy, S.; Abecasis, G.R.; et al. Reference-Based Phasing Using the Haplotype Reference Consortium Panel. *Nat. Genet.* **2016**, *48*, 1443–1448. [[CrossRef](#)] [[PubMed](#)]
38. Rubinacci, S.; Delaneau, O.; Marchini, J. Genotype Imputation Using the Positional Burrows Wheeler Transform. *PLoS Genet.* **2020**, *16*, e1009049. [[CrossRef](#)] [[PubMed](#)]
39. Cocca, M.; Barbieri, C.; Concas, M.P.; Robino, A.; Brumat, M.; Gandin, I.; Trudu, M.; Sala, C.F.; Vuckovic, D.; Girotto, G.; et al. A Bird’s-Eye View of Italian Genomic Variation through Whole-Genome Sequencing. *Eur. J. Hum. Genet.* **2020**, *28*, 435–444. [[CrossRef](#)] [[PubMed](#)]
40. Privé, F.; Arbel, J.; Vilhjálmsson, B.J. LDpred2: Better, Faster, Stronger. *Bioinformatics* **2021**, *36*, 5424–5431. [[CrossRef](#)]
41. Loh, P.-R.; Kichaev, G.; Gazal, S.; Schoech, A.P.; Price, A.L. Mixed-Model Association for Biobank-Scale Datasets. *Nat. Genet.* **2018**, *50*, 906–908. [[CrossRef](#)]
42. *Studio: Integrated Development Environment for R. Posit Software*, PBC: Boston, MA, USA, 2022.

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